

Programming and Post-Estimation

- Bootstrapping
- Monte Carlo
- Post-Estimation Simulation (Clarify)
- Extending Clarify to Other Models
 - Censored Probit Example

What is Bootstrapping?

- A computer-simulated nonparametric technique for making inferences about a population parameter based on sample statistics.
- If the sample is a good approximation of the population, the sampling distribution of interest can be estimated by generating a large number of new samples from the original.
- Useful when no analytic formula for the sampling distribution is available.

How do I do it?

$$SE_{\hat{E}_B} = \sqrt{\sum_{b=1}^B [s(x^B) - \sum_{b=1}^B s(x^b) / B]^2 / (B-1)}$$

1. Obtain a Sample from the population of interest. Call this $\mathbf{x} = (x_1, x_2, \dots, x_n)$.
2. Re-sample based on \mathbf{x} by randomly sampling with replacement from it.
3. Generate many such samples, x^1, x^2, \dots, x^B – each of length n .
4. Estimate the desired parameter in each sample, $s(x^1), s(x^2), \dots, s(x^B)$.
5. For instance the bootstrap estimate of the standard error is the standard deviation of the bootstrap replications.

Example: Standard Error of a Sample Mean

Canned in Stata



Intercooled Stata 8.2

File Edit Prefs Data Graphics Statistics User Window Help

Review
use "I:\general\PRISM Programming
sum mpg

Stata Results

tm
Statistics/Data Analysis 8.2 Copyright 1984-2003
StataCorp
4905 Lakeway Drive
College Station, Texas 77845 USA
800-STATA-PC http://www.stata.com
979-696-4600 stata@stata.com
979-696-4601 <fax>

48-student Stata for Windows (network) perpetual license:
Serial number: 1980523842
Licensed to: Ohio State University
Ohio State University

Notes:
1. </m# option or -set memory-> 1.00 MB allocated to data

. use "I:\general\PRISM Programming\auto.dta", clear
<1978 Automobile Data>

. sum mpg

Variable	Obs	Mean	Std. Dev.	Min	Max
mpg	74	21.2973	5.785503	12	41

Variables
Target: Command Window
make
price
mpg
rep78
headroom
trunk
weight
length
turn
displacement
gear_ratio
foreign

Stata Command

C:\DATA

Example: Standard Error of a Sample Mean Canned in Stata

The screenshot shows the Stata 8.2 interface with the following components:

- Review window:** Contains the command sequence:


```
use "I:\general\PRISM Programming
sum mpg
bootstrap "sum mpg" r(mean), reps(1
```
- Variables window:** Lists variables: make, price, mpg, rep78, headroom, trunk, weight, length, turn, displacement, gear_ratio, foreign.
- Stata Results window:**
 - Command: `sum mpg`
 - Statistic: `_bs_1 = r(mean)`
 - Warning: Since `sum` is not an estimation command or does not set `e(sample)`, `bootstrap` has no way to determine which observations are used in calculating the statistics and so assumes that all observations are used. This means no observations will be excluded from the resampling due to missing values or other reasons.
 - Text: If the assumption is not true, press Break, save the data, and drop the observations that are to be excluded. Be sure the dataset in memory contains only the relevant data.
 - Bootstrap statistics summary:

Variable	Reps	Observed	Bias	Std. Err.	[95% Conf. Interval]
_bs_1	1000	21.2973	.0145811	.6790806	19.96471 22.62988 (N) 20.05405 22.62162 (P) 20.08108 22.7027 (BC)
 - Note:
 - N = normal
 - P = percentile
 - BC = bias-corrected
- Stata Command window:** Empty.

$$\approx 21.2973 \pm 1.96 * .6790806$$

$$\bar{x}^B - \bar{x}$$

Example: Difference of Medians Test

The screenshot shows the Intercooled Stata 8.2 interface. The **Review** window on the left contains the following Stata program:

```
use "I:\general\PRISM Programming  
sum mpg  
bootstrap "sum mpg" r(mean), reps(1  
do "C:\DOCUMENT~1\ADMINI~1\LO  
bootstrap "mymedian" r(diff), reps(10  
sum length, detail  
return list  
sum length, detail  
return list  
sum length, detail  
return list  
sum length, detail  
return list
```

The **Variables** window on the left lists the following variables: make, price, mpg, rep78, headroom, trunk, weight, length, turn, displacement, gear_ratio, foreign.

The **Stata Results** window displays the following output:

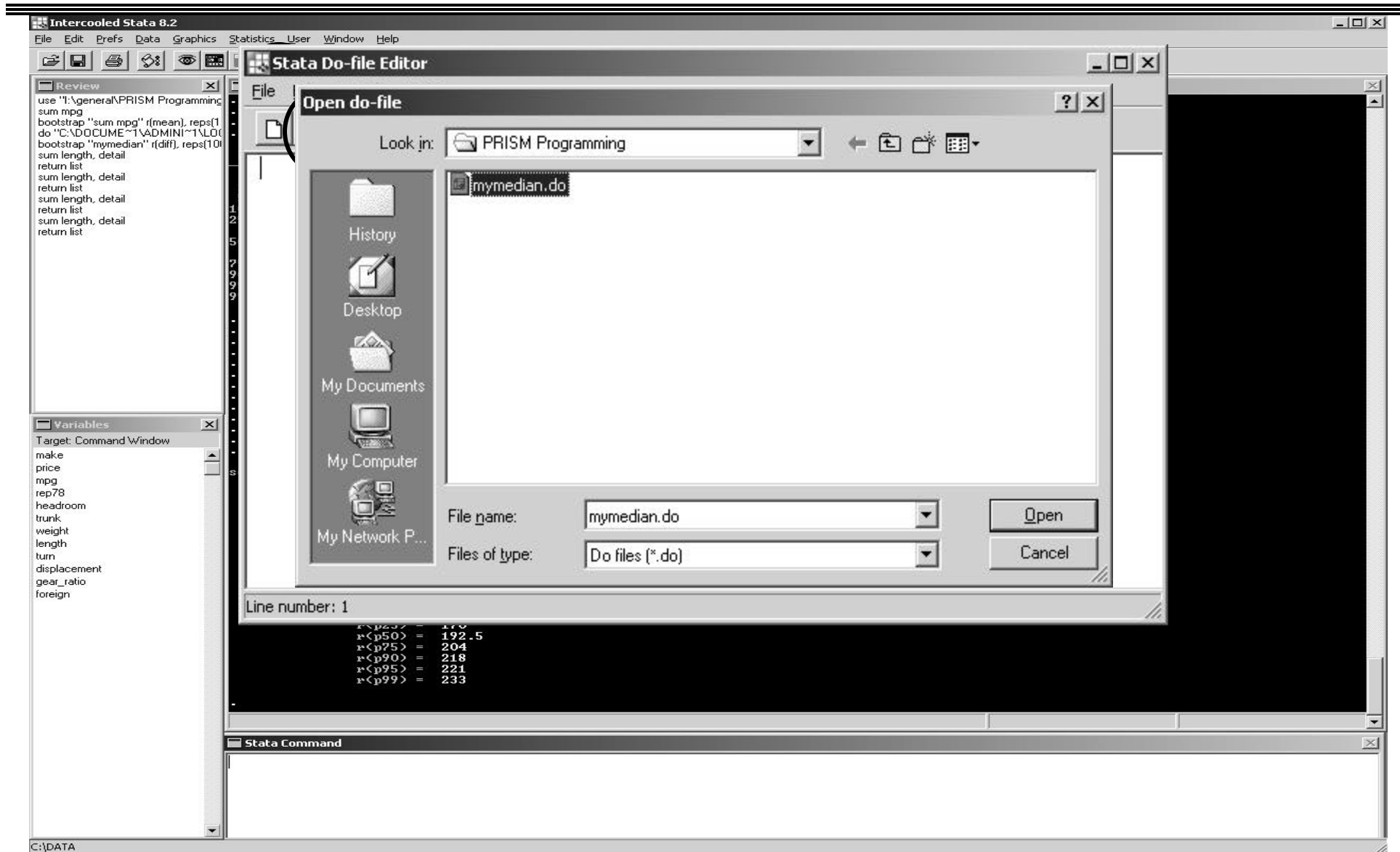
```
. sum length, detail  
  
Length (in.)  
  
Percentiles      Smallest  
1%              142      142  
5%              154      147  
10%             157      149  
25%             170      154  
  
50%             192.5  
75%             204  
90%             218  
95%             221  
99%             233  
  
Largest  
                221  
                222  
                230  
                233  
  
Obs              74  
Sum of Wgt.      74  
  
Mean             187.9324  
Std. Dev.        22.26634  
Variance         495.7899  
Skewness         -.0409746  
Kurtosis         2.04156
```

The **return list** command shows the following scalars:

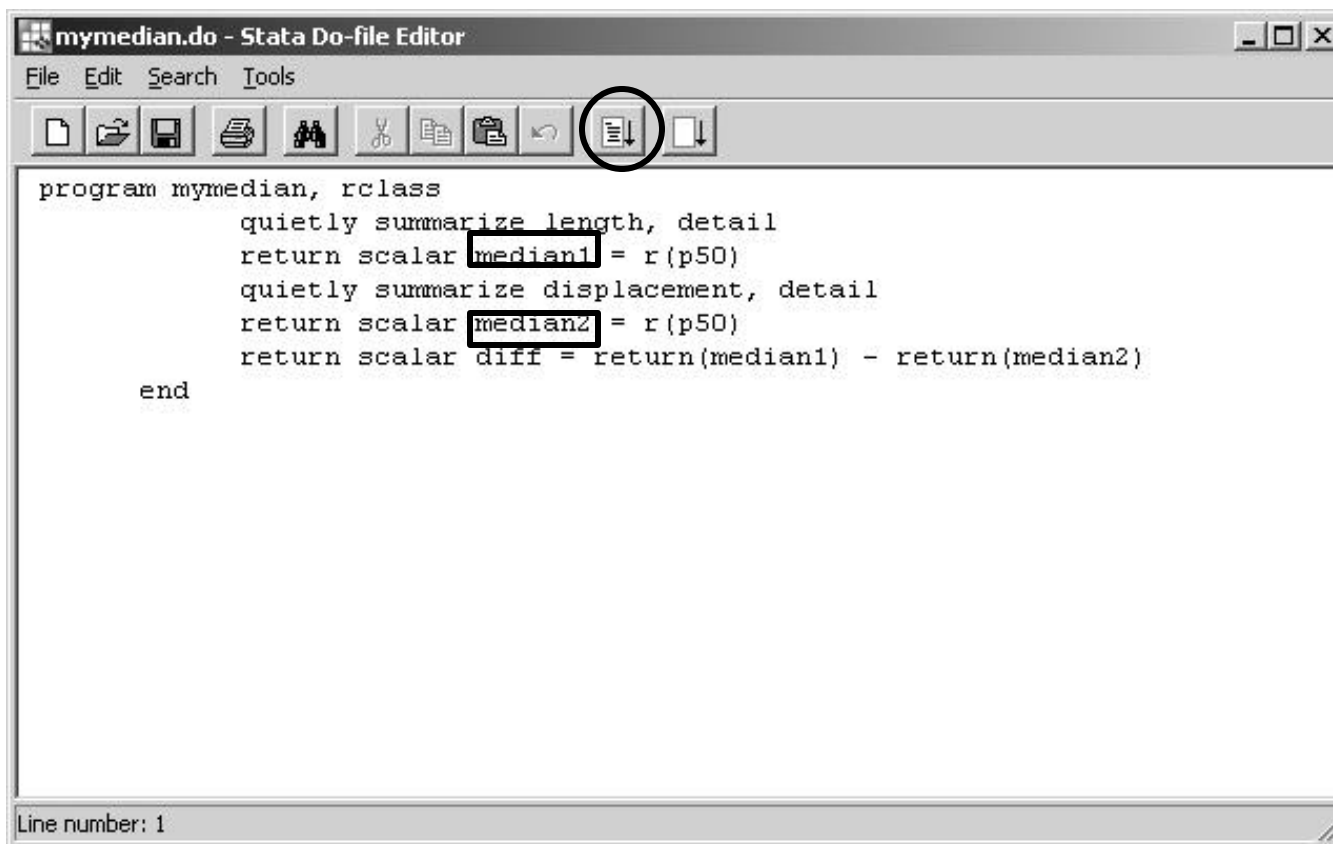
```
. return list  
scalars:  
r(N) = 74  
r(sum_w) = 74  
r(mean) = 187.9324324324324  
r(Uar) = 495.7898926323584  
r(sd) = 22.26633990202158  
r(skewness) = -.0409745547984551  
r(kurtosis) = 2.041559782872945  
r(sum) = 13907  
r(min) = 142  
r(max) = 233  
r(p1) = 142  
r(p5) = 154  
r(p10) = 157  
r(p25) = 170  
r(p50) = 192.5  
r(p75) = 204  
r(p90) = 218  
r(p95) = 221  
r(p99) = 233
```

The **Stata Command** window is empty.

Example: Difference of Medians Test



Example: Difference of Medians Test



The screenshot shows the Stata Do-file Editor window titled "mymedian.do - Stata Do-file Editor". The window has a menu bar with "File", "Edit", "Search", and "Tools". Below the menu bar is a toolbar with various icons. The main text area contains the following Stata code:

```
program mymedian, rclass
    quietly summarize length, detail
    return scalar median1 = r(p50)
    quietly summarize displacement, detail
    return scalar median2 = r(p50)
    return scalar diff = return(median1) - return(median2)
end
```

The status bar at the bottom left indicates "Line number: 1".

Example: Difference of Medians Test

Intercooled Stata 8.2

File Edit Prefs Data Graphics Statistics User Window Help

Review

use "I:\general\PRISM Programming
do "C:\DOCUME~1\ADMINI~1\LOCALS~1\Temp\STD01000000.tmp"
bootstrap "mymedian" r(diff), reps(1000)

Stata Results

```
do "C:\DOCUME~1\ADMINI~1\LOCALS~1\Temp\STD01000000.tmp"

. program mymedian, rclass
1.         quietly summarize length, detail
2.         return scalar median1 = r(p50)
3.         quietly summarize displacement, detail
4.         return scalar median2 = r(p50)
5.         return scalar diff = return(median1) - return(median2)
6.         end

. end of do-file

. bootstrap "mymedian" r(diff), reps(1000)

command:      mymedian
statistic:    _bs_1      = r(diff)

Warning: Since mymedian is not an estimation command or does not set e(sample), bootstrap has no way to determine which
observations are used in calculating the statistics and so assumes that all observations are used. This means no
observations will be excluded from the resampling due to missing values or other reasons.

If the assumption is not true, press Break, save the data, and drop the observations that are to be excluded. Be sure the
dataset in memory contains only the relevant data.

Bootstrap statistics              Number of obs   =       74
                                Replications    =     1000

+-----+-----+-----+-----+-----+-----+-----+-----+
Variable | Reps | Observed | Bias | Std. Err. | [95% Conf. Interval] |
+-----+-----+-----+-----+-----+-----+-----+
       _bs_1 | 1000 |      -3.5 |  5.765 | 24.48127 | -51.54061   44.54061 |
              |      |          |      |          |          -36   38.5 |
              |      |          |      |          |          -37.5  38  <BC> |
+-----+-----+-----+-----+-----+-----+-----+

Note:  N = normal
       P = percentile
       BC = bias-corrected
```

Variables

Target: Command Window

make
price
mpg
rep78
headroom
trunk
weight
length
turn
displacement
gear_ratio
foreign

Stata Command

C:\DATA

What is Monte Carlo Simulation?

- Uses the observation of random samples from known populations of simulated data to track the behavior of a statistic.
- If the sampling distribution of a statistic is the density function of values it could take on in a given population, then its estimate is the relative frequency distribution of the values that were actually observed in many samples drawn from that population.
- Since we can generate the population to have any characteristics we wish, Monte Carlos are very flexible.

How do I do it?

1. Determine what the “population” is.
2. Sample from the “population”
3. Calculate the estimator of interest \hat{q} . Save this value.
4. Repeat steps 2 and 3 many times.
5. Construct the frequency distribution of the \hat{q} values. This is the Monte Carlo estimate of the sampling distribution of q under the conditions you specified for the population.

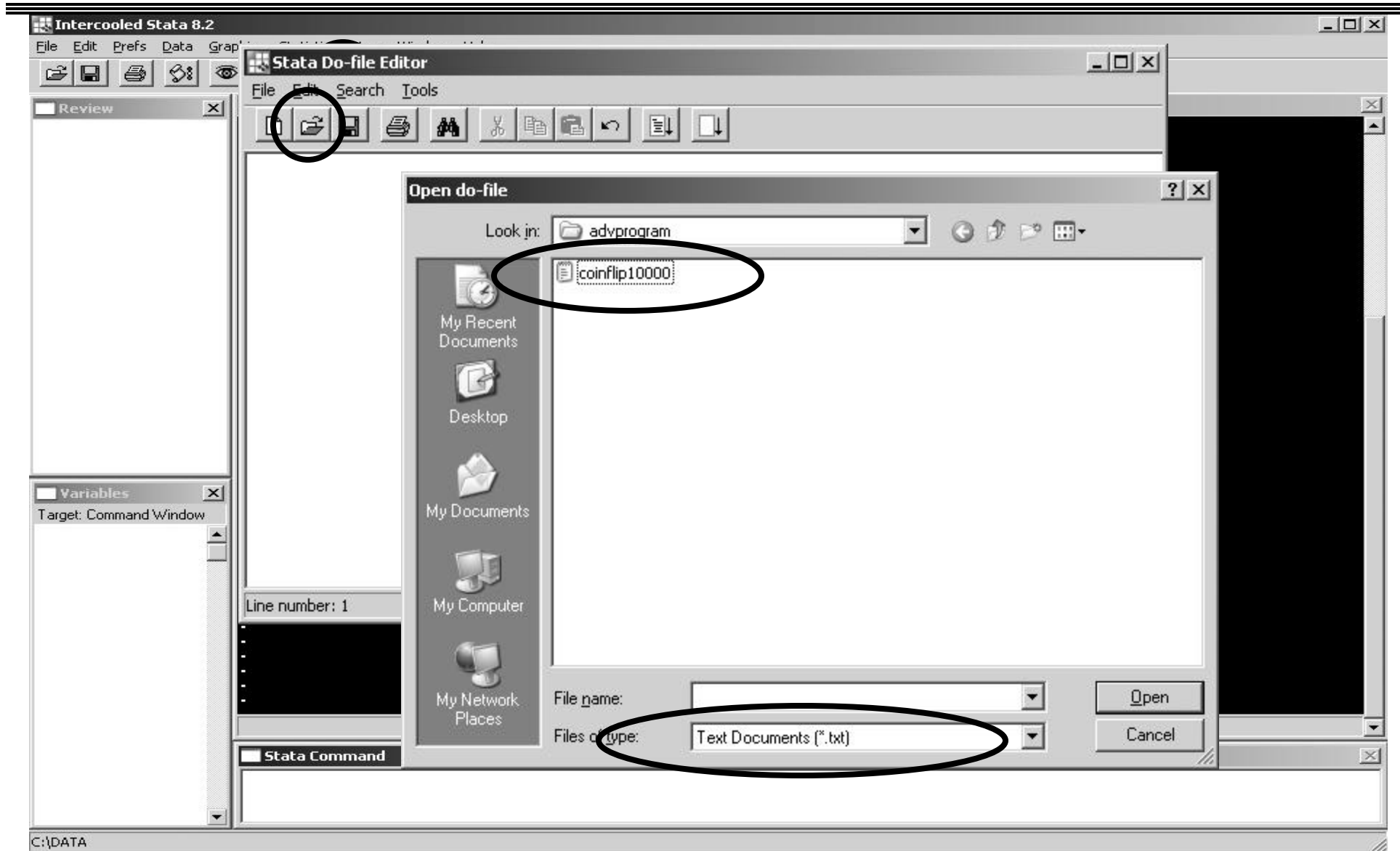
An Example: Coin Toss

- If you toss a fair coin 10 times, what is the probability of obtaining exactly 3 heads?
- By the binomial probability distribution:

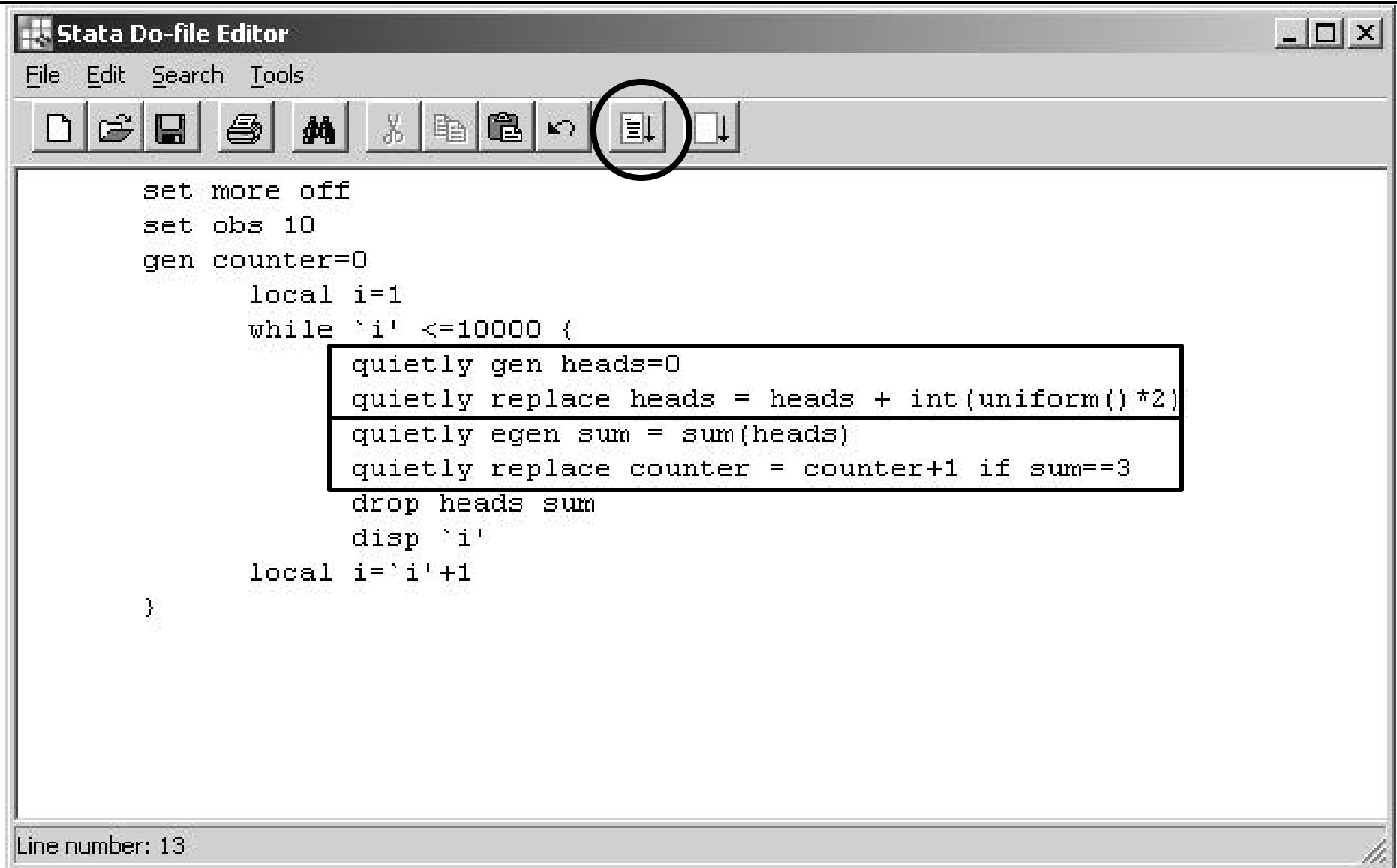
$$= \frac{10!}{3!(10-3)!} 0.5^3 * 0.5^7 \approx 0.1172$$

However, if we didn't know how to use the binomial distribution, we could toss a fair coin 10 times in a repeated number of trials and simulate this probability.

Coin Toss Monte Carlo



Coin Toss Monte Carlo



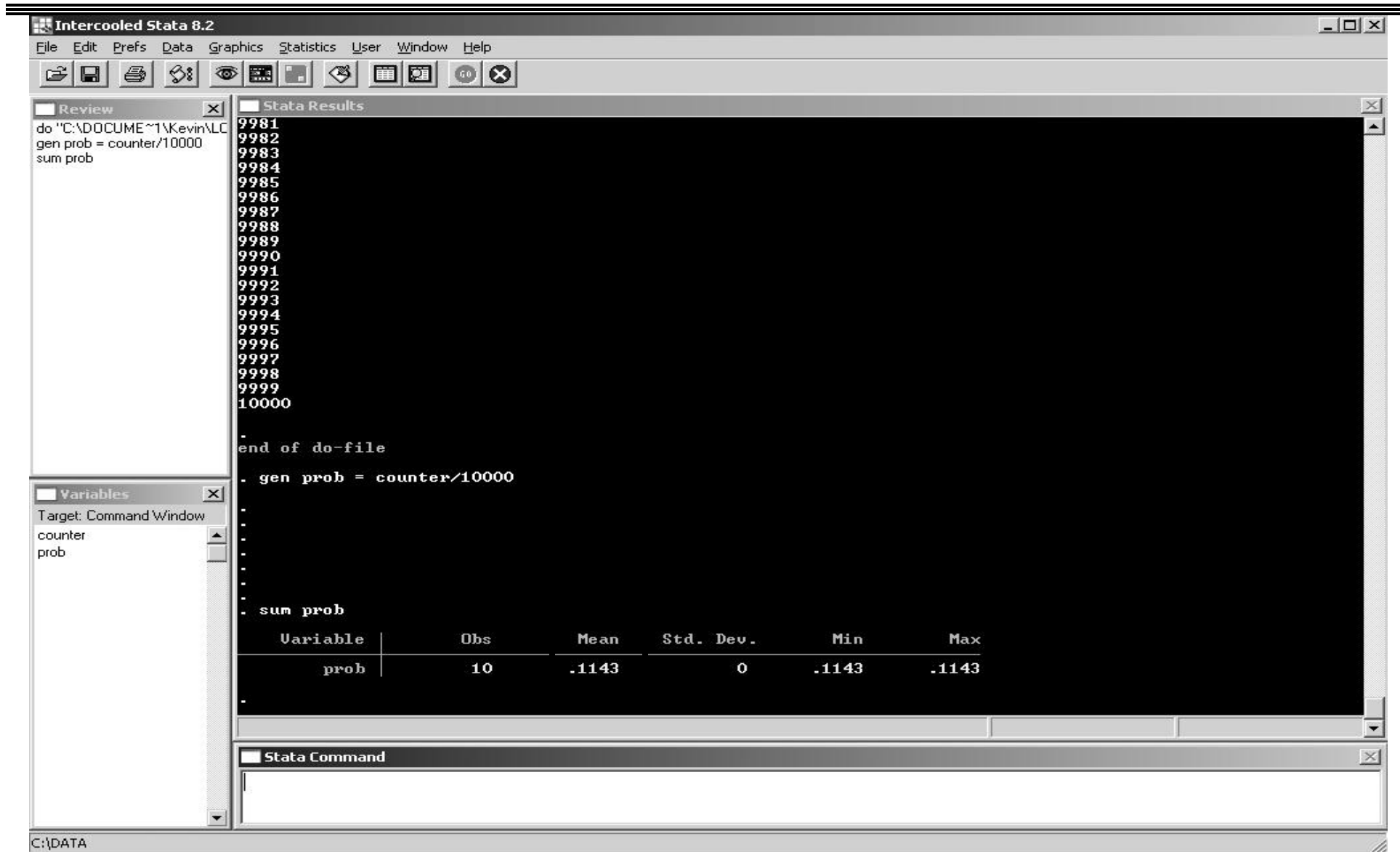
Stata Do-file Editor

File Edit Search Tools

set more off
set obs 10
gen counter=0
 local i=1
 while `i' <=10000 {
 quietly gen heads=0
 quietly replace heads = heads + int(uniform()*2)
 quietly egen sum = sum(heads)
 quietly replace counter = counter+1 if sum==3
 drop heads sum
 disp `i'
 local i=`i'+1
}

Line number: 13

Coin Toss Monte Carlo



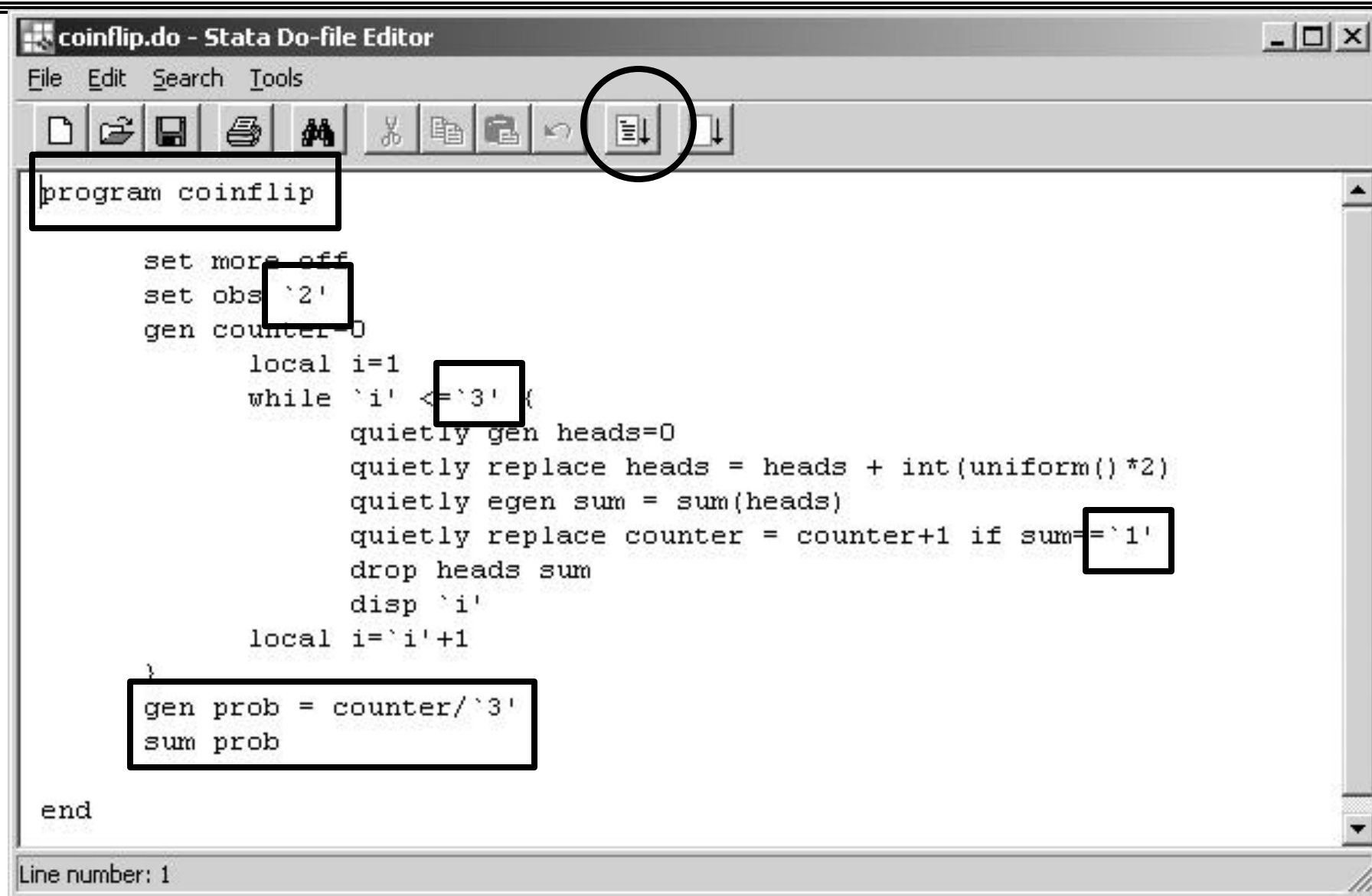
Programming the Coin Toss Monte Carlo

- Allows us to automate the experiment, controlling:
 - ✓ The number of tosses
(e.g. 10, or something else)
 - ✓ The number of trials
(e.g. 10,000 or something else)
 - ✓ The number of successes
(e.g. 3 or something else)

Coin Toss Program



Coin Toss Program



```
coinflip.do - Stata Do-file Editor
File Edit Search Tools

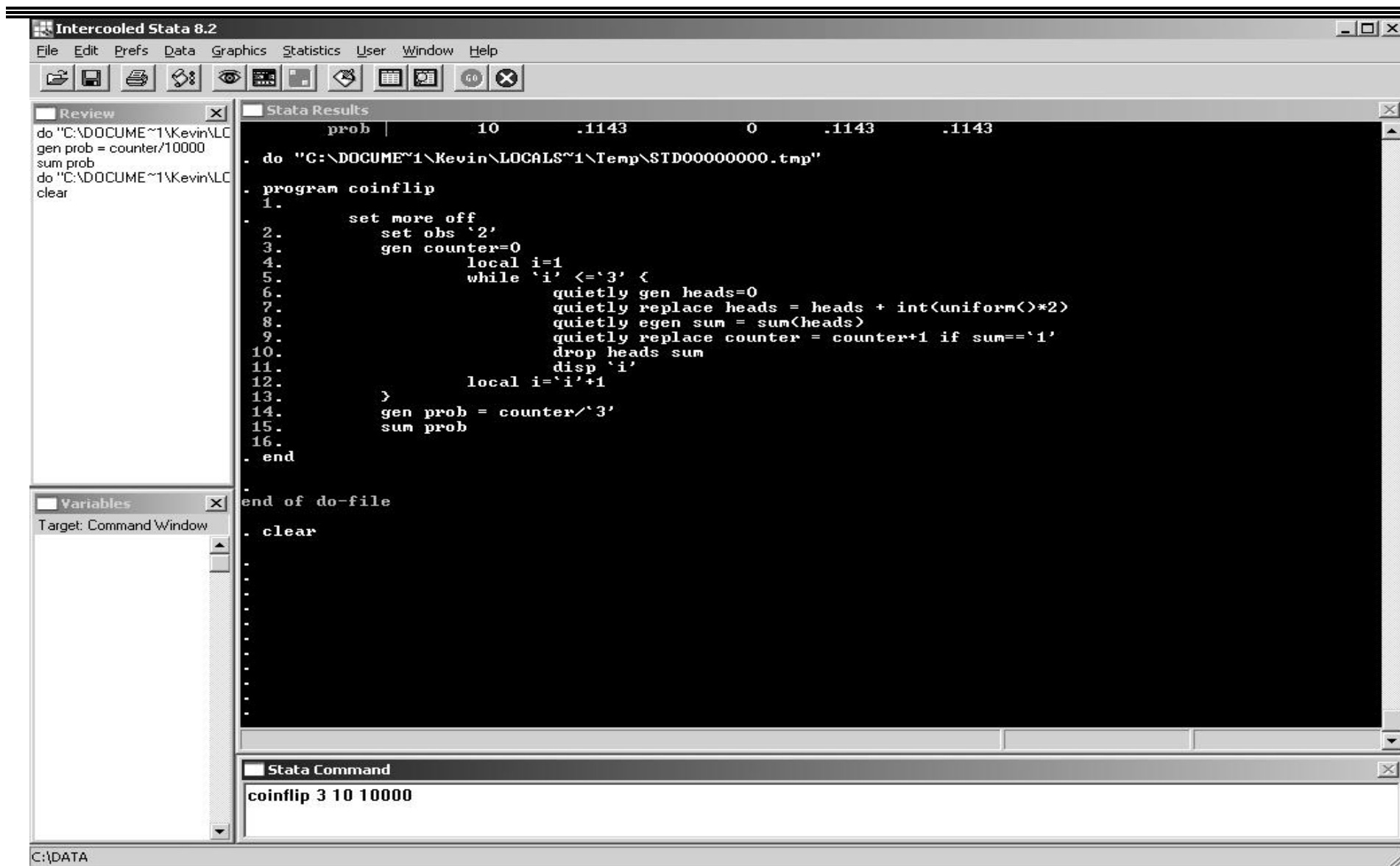
program coinflip

set more off
set obs `2'
gen counter=0
    local i=1
    while `i' <= `3' {
        quietly gen heads=0
        quietly replace heads = heads + int(uniform()*2)
        quietly egen sum = sum(heads)
        quietly replace counter = counter+1 if sum==`1'
        drop heads sum
        disp `i'
        local i=`i'+1
    }
    gen prob = counter/`3'
    sum prob

end

Line number: 1
```

Coin Toss Program



Coin Toss Program

The screenshot displays the Intercooled Stata 8.2 interface. The main window is divided into several panes:

- Review:** Contains the following commands:

```
do "C:\DOCUME~1\Kevin\LC
gen prob = counter/10000
sum prob
do "C:\DOCUME~1\Kevin\LC
clear
coinflip 3 10 10000
```
- Variables:** Shows the target window and lists the variables `counter` and `prob`.
- Stata Results:** Displays the output of the commands. It lists observation numbers from 9967 to 10000. Below this, a summary table is shown for the variable `prob`.
- Stata Command:** An empty text area for entering commands.

The summary table in the Stata Results pane is as follows:

Variable	Obs	Mean	Std. Dev.	Min	Max
prob	10	.1165	0	.1165	.1165

The status bar at the bottom left indicates the current directory is `C:\DATA`.

Another Monte Carlo Application: The Logic of Post Estimation Simulation

So, you estimate a model... and you want to say something *substantive* with quantities of interest:

Predicted or Expected Values of DV = $X_m \hat{b}$

First Differences = $X_{+s} \hat{b} - X_m \hat{b}$

The problem is that our \hat{b}_s are uncertain!

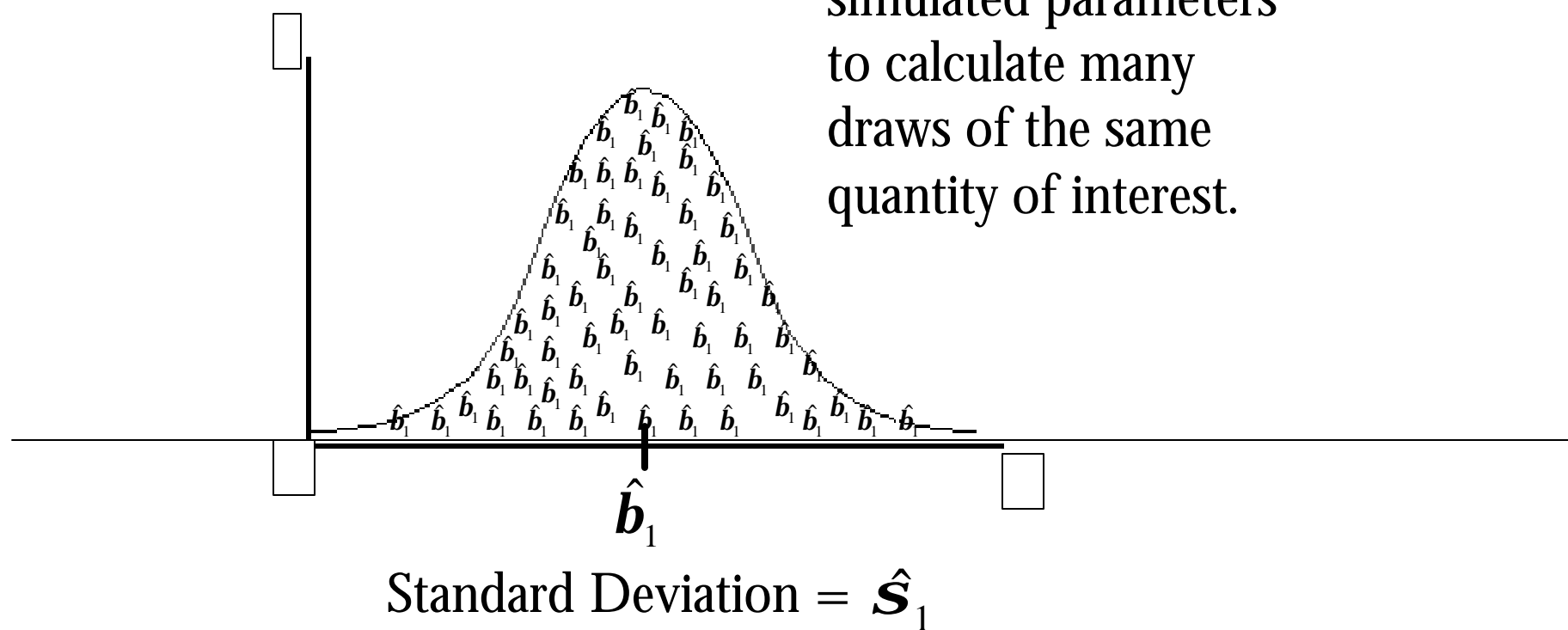
The solution is we know how uncertain.

$$\frac{\hat{b}_1}{\hat{s}_1}$$

Monte Carlo Simulation of Parameters

In order to capture the uncertainty, we draw simulated $\hat{\mathbf{b}}$ s from the multivariate* normal distribution.

Then we use these simulated parameters to calculate many draws of the same quantity of interest.



Simulating Quantities of Interest

In practice... $Y_i \sim f(\mathbf{q}_i, \mathbf{a}), \quad \mathbf{q}_i = g(X_i, \mathbf{b})$
 $Y_i \sim N(\mathbf{m}_i, \mathbf{S}^2), \quad \mathbf{m}_i = g(X_i, \mathbf{b}) = \mathbf{b}_0 + X_{i1} \mathbf{b}_1 + X_{i2} \mathbf{b}_2 + \dots$

$$\hat{\mathbf{g}} = \begin{bmatrix} \hat{\mathbf{b}}_1 \\ \hat{\mathbf{b}}_2 \\ \vdots \\ \hat{\mathbf{a}}_1 \end{bmatrix} \quad \hat{V}(\hat{\mathbf{g}}) = \begin{bmatrix} v_{\hat{\mathbf{b}}_1 \hat{\mathbf{b}}_1} & v_{\hat{\mathbf{b}}_1 \hat{\mathbf{b}}_2} & \cdots & v_{\hat{\mathbf{b}}_1 \hat{\mathbf{a}}} \\ v_{\hat{\mathbf{b}}_2 \hat{\mathbf{b}}_1} & v_{\hat{\mathbf{b}}_2 \hat{\mathbf{b}}_2} & \cdots & v_{\hat{\mathbf{b}}_2 \hat{\mathbf{a}}} \\ \vdots & \vdots & \ddots & \vdots \\ v_{\hat{\mathbf{a}} \hat{\mathbf{b}}_1} & v_{\hat{\mathbf{a}} \hat{\mathbf{b}}_2} & \cdots & v_{\hat{\mathbf{a}} \hat{\mathbf{a}}} \end{bmatrix}$$

we simulate parameters with M draws from the multivariate normal distribution... $\tilde{\mathbf{g}} \sim N(\hat{\mathbf{g}}, \hat{V})$

$$\begin{bmatrix} \tilde{\mathbf{b}}_{11} \\ \tilde{\mathbf{b}}_{21} \\ \vdots \\ \tilde{\mathbf{a}}_1 \end{bmatrix} \begin{bmatrix} \tilde{\mathbf{b}}_{12} \\ \tilde{\mathbf{b}}_{22} \\ \vdots \\ \tilde{\mathbf{a}}_2 \end{bmatrix} \dots \begin{bmatrix} \tilde{\mathbf{b}}_{1M} \\ \tilde{\mathbf{b}}_{2M} \\ \vdots \\ \tilde{\mathbf{a}}_M \end{bmatrix}$$

- Choose a starting scenario, X_c .
- Draw one value of $\tilde{\mathbf{g}}$, and compute $\tilde{\mathbf{q}}_c = g(X_c, \tilde{\mathbf{b}})$.
- Simulate the outcome \tilde{Y}_c , by taking a random draw from $f(\tilde{\mathbf{q}}_c, \tilde{\mathbf{a}})$.
- Repeat M times to get the distribution of Y_c .

Clarify (King et. al. *AJPS* 1999)

- *estsimp* – estimates the model and simulates the parameters
 - This command **must** precede your regression command
 - e.g.: *estsimp logit y x1 x2 x3 x4*
 - This will save simulated β s to your dataset!
- *setx* – sets the values for the IVs (the Xs)
 - Used after model estimation to set values of the Xs
 - e.g.: *setx x1 mean x2 p20 x3 .4 x4[16], nocwdel*
 - functions = mean|median|min|max|p#|math|#|‘macro’|varname[#]
 - reset values by re-issuing the command, e.g.: *setx x1 median*
- *simqi* – simulates the quantities of interest
 - Automates the simulation of quantities of interest for the X values you just set.
 - e.g.: *simqi, prval(1)*
 - e.g.: *simqi, fd(prval(1)) changex(x4 p25 p75)*

You Can Use Clarify, but you Don't have to.

Models Currently Supported by Clarify

regress	mlogit
logit	poisson
probit	nbreg
ologit	sureg
oprobit	weibull

But, you really don't need Clarify to do this, so you can simulate quantities of interest for *any* model!

- ✓ Easy to simulate parameters because Stata saves them after estimation!
- ✓ Program the correct link function yourself!

An Example: The Censored Probit Model

Selection Equation:

$$y1_j = z_j \mathbf{g} + u_{2j}$$

Outcome Equation:

$$y2_j = x_j \mathbf{b} + u_{1j}$$

Where:

$$u_1 \sim N(0,1)$$

$$u_2 \sim N(0,1)$$

$$\text{corr}(u_1, u_2) = \mathbf{r}$$

/*Programming Step One:
Estimate Model*/

***heckprob y2 x1 x2 x3 x4,
sel(y1 = z1 z2 z3 z4)
robust***

An Example:

The Censored Probit Model

Simulate the model parameters by drawing from the multivariate normal distribution.

Note: there are 11 – 4 Xs, 4 Zs, 2 constants, and ρ (the correlation between the errors).

/*Programming Step Two:
Draw \tilde{b} from multivariate normal, mean \hat{b} and
Covariance Matrix $\hat{\Sigma}$.*/

matrix params = e(b)

matrix P = e(V)

***drawnorm b1-b11,
means(params) cov(P)
double***

An Example: The Censored Probit Model

Stata estimates the
hyperbolic arctangent
of ρ , so we must
simulate to get the
actual ρ .

/*Programming Step Three:
Generate Simulated Rho*/
***gen simrho = (exp(2*b11)-
1)/(exp(2*b11)+1)***

An Example:

The Censored Probit Model

Initiate a looping structure to generate m (in this case 1,000) simulated first differences for the effect of x_1 on y_2 comparing when x_1 is at its mean (the base model) to when x_1 is at a value two standard deviations (denoted $_m2sd$) below its mean.

```
/*ProgramStep Four: The Loop*/
```

```
local i =1
```

```
/*A. Generate variables that will be used to fill  
in a cell of Substantive Table*/
```

```
generate base_y2=.
```

```
generate x1_m2sd=.
```

```
while `i' <=1000 {
```

```
/*B. Generate  $z\gamma$  for the selection equation.*/
```

```
quietly generate select = b6[`i'] +  
(b7[`i']*z1) + (b8[`i']*z2) +  
(b9[`i']*z3) + (b10[`i']*z4)
```

```
/*C. Generate  $x_b$  for the outcome equation.*/
```

```
quietly generate outcome =  
b1[`i'] + (b2[`i']*x1) +  
(b3[`i']*x2) + (b4[`i']*x3) +  
(b5[`i']*x4)
```

An Example:

The Censored Probit Model

This is the meat of the simulation. The first three commands generate the probability of being selected and experiencing the outcome (p_{11}) for the base model. For the censored probit, this probability is (Greene 2000, 857):

$$\Phi_2[b'x, g'z, r]$$

```
/* Generate first difference */  
quietly generate p_11 =  
    binorm(outcome,select,simrho)  
quietly summarize p_11, meanonly  
quietly replace base y2=r(mean) in `i'  
quietly generate x1_m2sd=outcome -  
    (b1[`i']*x1) + (b1[`i']* -0.2)  
quietly generate p11_x1_m2sd  
    =binorm(x1_m2sd,select,simrho)  
quietly summarize p_x1_m2sd,  
    meanonly  
quietly replace x1_m2sd=r(mean) in `i'
```

An Example:

The Censored Probit Model

To get the other 999 we drop the three variables we just generated and repeat the loop until ``i' = 1,000`.

When we're done with the 1,000 simulations, we can use the `centile` command to get the relevant distributions. To do this for each variable in the model, we would embed this loop within a larger looping structure.

```
/*Step 5: Do the Loop Again*/  
drop select outcome p_11  
x1_m2sd p11_x1_m2sd  
  
disp `i'  
local i=`i'+1  
}  
  
centile base_y2 x1_m2sd,  
centile(2.5 50 97.5)
```